



# Time scales of adaptation to context in horizontal sound localization

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# **ABSTRACT:**

Psychophysical experiments explored how the repeated presentation of a context, consisting of an adaptor and a target, induces plasticity in the localization of an identical target presented alone on interleaved trials. The plasticity, and its time course, was examined both in a classroom and in an anechoic chamber. Adaptors and targets were 2 ms noise clicks and listeners were tasked with localizing the targets while ignoring the adaptors (when present). The context was either simple, consisting of a single-click adaptor and a target, or complex, containing either a single-click or an eight-click adaptor that varied from trial to trial. The adaptor was presented either from a frontal or a lateral location, fixed within a run. The presence of context caused responses to the isolated targets to be displaced up to 14° away from the adaptor location. This effect was stronger and slower if the context was complex, growing over the 5 min duration of the runs. Additionally, the simple context buildup had a slower onset in the classroom. Overall, the results illustrate that sound localization is subject to slow adaptive processes that depend on the spatial and temporal structure of the context and on the level of reverberation in the environment. © 2023 Acoustical Society of America. https://doi.org/10.1121/10.0021304

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#### I. INTRODUCTION

Auditory spatial perception is highly adaptive (Carlile, 2014; King et al., 2000). Changes in horizontal sound localization can be induced by visual stimulation (Recanzone, 1998), feedback training (Klingel et al., 2021; Shinn-Cunningham et al., 1998), a change in the acoustic environment (Shinn-Cunningham et al., 2005), alterations in the mapping between acoustic cues and source locations (Kumpik et al., 2010; Trapeau and Schoenwiesner, 2018; van Wanrooij and van Opstal, 2007), or by other stimuli presented either simultaneously with the target (Braasch and Hartung, 2002), or preceding the target (Kopčo et al., 2010). The adaptation induced by preceding stimulation has been observed for adaptor stimuli presented over tens of seconds and minutes, e.g., in the auditory localization aftereffect (Carlile et al., 2001; Phillips and Hall, 2005; Thurlow and Jack, 1973), or in the precedence buildup induced by repeated presentation of "lead-lag" stimulus pairs (Djelani and Blauert, 2001; Freyman et al., 1991). Studies of auditory localization aftereffects typically used a long continuous adaptor immediately followed by a target (Carlile et al., 2001; Thurlow and Jack, 1973), or even overlapping with the target (Canévet and Meunier, 1996). They observed a repulsion by the adaptor, i.e., biases in the perceived target locations away from the adaptor location. Other studies have found similar effects for briefer adaptors that change location from trial to trial (e.g., Laback, 2023), calling into

question the need for very prolonged stimulation to observe localization aftereffects. Studying such effects is important because everyday listening almost always occurs in complex multi-source environments and how preceding stimulation on different time scales influences perception is not well understood. Here, we examine an adaptive effect qualitatively similar to the localization aftereffect but induced by the trial-to-trial acoustic "context" in which target sounds are presented. In our experiments, the target is a 2 ms broadband noise burst (referred to here as a "click") (Kopčo et al., 2007). On some trials, it is immediately preceded by an identical adaptor click (or clicks), and on other trials, it is presented in isolation. Of interest here are localization biases for the target-only trials that are induced when those trials are randomly interleaved with adaptor-target trials. This effect, called contextual plasticity (CP), was observed in our previous work as repulsive biases of up to  $10^{\circ}$  in localization of the single-click targets.

Several different mechanisms have been proposed as underlying localization biases. First, some adaptation or fatiguing of the peripheral neural representation due to prolonged stimulation is often assumed (Carlile *et al.*, 2001; Flugel, 1920). Second, a rebalancing of the putative hemispheric channels subserving spatial processing in humans has been proposed (Dingle *et al.*, 2012; Phillips and Hall, 2005). Third, recent models based on known physiology of subcortical binaural circuits suggest that adaptation in response to the preceding context causes a rescaling of the spatial representation with the goal of increasing perceptual spatial separability of frequently presented sounds at the

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cost of inducing localization biases (Dahmen *et al.*, 2010; Lingner *et al.*, 2018; Maddox *et al.*, 2014). Finally, an active centrally driven suppression of reverberation has been proposed for the precedence buildup, a potentially related phenomenon (Clifton *et al.*, 2002).

The current study is the fourth in a series that examines CP. The original study (Kopčo *et al.*, 2007) reported CP as an unexpected effect observed both in anechoic and reverberant rooms. Kopčo *et al.* (2015) showed that the effect is driven by adaptation in auditory perceptual representations as opposed to motor response–related representations, as it was observed for various response methods and with or without visual inputs. Hládek *et al.* (2017) showed that the strength of CP depends on the number of adaptor clicks and their similarity to the target. The goal of this fourth study is to examine how complexity of the context affects CP and to present a detailed analysis of the temporal profile of CP.

Our analysis is based on data from two experiments using an identical design: one performed in a small classroom (experiment 1) and one performed in an anechoic chamber (experiment 2). While the experiments were originally designed to examine the fast adaptation effects of the immediately preceding adaptors on time scales shorter than 0.5 s (Kopčo et al., 2007; Kopčo et al., 2017), the current study only focuses on the slower effects related to CP (some of which were reported in the previous studies without detailed analysis). In the experiments, CP was induced by context trials in which the adaptor was located either in front of, or to the side of, the listener [Fig. 1(A)], in one of two stimulus conditions [Fig. 1(B)]: in the simple context condition, the adaptor always contained one click (Kopčo et al., 2007), while in the *complex context* condition, the adaptor was either a single click or a train of eight clicks, varying from trial to trial (Kopčo et al., 2017).



FIG. 1. Experimental setup and stimuli. (A) Arrangement of the loudspeaker array (shown here on the subject's right-hand side). The adaptor (loudspeaker in gray color) was in the frontal position for half of the runs and in the lateral position for the other half. (B) Temporal structure of the target-only and adaptor-target stimuli, with adaptor in gray and target in white. Inter-stimulus interval, the time interval between the final adaptor click onset to the target click onset, ranged from 25 to 400 ms. Exp. 1, experiment 1; Exp. 2, experiment 2.



We addressed several questions related to the time course of CP. First, while we expected that CP would be stronger in the complex context condition as the average number of context clicks is higher in this condition (Hládek et al., 2017), we tested the hypothesis that it might also be somewhat unstable as the context varies from trial to trial and as a result, take longer to asymptote (as observed in other auditory domains, e.g., Vilfan and Duke, 2003). Second, we hypothesized that CP may be weaker and/or slower in the reverberant environment, as reverberation tends to make the spatiotemporal distribution of the received stimuli more uniform, which may reduce the strength of adaptation. For example, in anechoic space, a lateral adaptor delivers energy exclusively from its lateral azimuth. In reverberation, however, it will be accompanied by multiple reflections with distributed azimuths and on aggregate, the energy will be less lateralized (Shinn-Cunningham et al., 2005) and may cause less spatial adaptation. Finally, we examined whether CP has a fast component on the time scale of seconds, observable when the context is varying from trial to trial.

## **II. METHODS**

The data described here were collected as part of two experiments previously reported in (Kopčo *et al.*, 2007; Kopčo *et al.*, 2017). While details on the subjects, environments, and stimuli can be found in those publications, they are briefly described again here.

#### A. Subjects

Seven listeners (three females), with ages ranging from 23 to 32 years, participated in experiment 1 (classroom), and four of these listeners also participated in experiment 2 (anechoic room). All listeners had normal hearing as confirmed by audiometric screening (all thresholds within 15 dB hearing level) and gave informed consent as approved by the Institutional Review Board. Of the subjects who participated in both experiments, three started with experiment 1, one with experiment 2, an imbalance that was not expected to strongly influence the results.

#### B. Setup and listening environment

Experiment 1 was conducted in an empty, quiet rectangular reverberant room. The reverberation times in octave bands centered at 500, 1000, 2000, and 4000 Hz were 613, 508, 512, and 478 ms, respectively. The background noise level was 39 dBA. Experiment 2 was conducted in an anechoic room. Nine loudspeakers (Bose Acoustimass, Bose, Framingham, MA) were positioned on an arc with diameter of 1.2 m spanning 90°. The listener was seated approximately in the center of either room with his/her head held stable by a headrest. He/she sat in the center of the arc and faced either the left-most loudspeaker [so that the targets occurred on his/her right; see Fig. 1(A)] or the rightmost loudspeaker [setup mirror-flipped compared to Fig. 1(A)]. In the following, 0° azimuth always represents the JASA

location directly ahead of the listener, and 90° is the location of the left- or right-most speaker (depending on the listener orientation). Digital stimuli were generated by a TDT System 3 audio interface and passed through power amplifiers (Crown D-75A, Crown Audio, Elkhart, IN) to the loudspeakers. The listeners kept their eyes closed during experimental runs and held a pointer in one hand to indicate the perceived direction of each target. A Polhemus FasTrak (Colchester, VT) electromagnetic tracker was used to measure the location of the listener's head and the listener's responses (for details, see Kopčo *et al.*, 2007).

#### C. Stimuli and procedure

The target was a single 2 ms frozen noise burst (click) presented at 67 dBA [Fig. 1(B)]. An identical click was used for the adaptor in the one-click context trials. Eight such clicks presented at the rate of 10/s (T = 100 ms) made up the adaptor in the eight-click context trials. Within a run, the context was either simple or complex. In the simple context runs, only the one-click contexts were used, the ratio of contextual to target-only trials was 5:1, and the adaptor-target inter-stimulus interval, measured from the adaptor click onset to the target click onset, was 25, 50, 100, 200, or 400 ms. In the complex context runs, the ratio of eight-click context to one-click context to target-only trials was 2:2:1 and the inter-stimulus interval was 50 or 200 ms. On each trial, the target location was randomly selected from one of the seven central loudspeakers (spanning approximately 11°-79° azimuth), while the adaptor, if any, was played from a loudspeaker that was fixed within a run. Every combination of the six (simple context) or five (complex context) trial types and seven target locations was presented four times in random order within a run, resulting in 168 trials in the simple context runs and 140 trials in the complex context runs. The subject changed his/her orientation after each run to face either the left-most loudspeaker or the right-most loudspeaker by rotating his/her whole body.

Experiments 1 and 2 each comprised eight sessions four for the simple context and four for the complex context. Each session, which took approximately 30 min, contained four randomly ordered runs—one for each combination of subject orientation (facing the left-most speaker, facing the right-most speaker) and context adaptor location (frontal, lateral). The total duration of a run was relatively consistent, with across-subject means and standard deviations of  $5.3 \pm 0.6$  min (experiment 1, simple context),  $5.1 \pm 0.7$  min (experiment 2, simple context),  $5.6 \pm 0.6$  min (experiment 1, complex context), and  $5.3 \pm 0.5$  min (experiment 2, complex context).

# D. Data analysis

The current analyses focus exclusively on data from the target-only trials (Kopčo *et al.*, 2007; Kopčo *et al.*, 2017). There were only small differences between the datasets collected with the two subject orientations, and thus, the data were collapsed across the orientations and sessions and

analyzed as if the subject always faced the left-most loudspeaker. Since only a subset of the experiment 1 subjects participated in experiment 2, data are also presented for this subset of four subjects in experiment 1, to allow a direct comparison of the effect of room across the subjects. To analyze the temporal profile of CP, data from each run were divided into four subruns, as each run contained four repetitions of each stimulus combination, presented in a pseudorandom order such that any combination was repeated only after all other combinations were presented. All reported statistical analyses were performed as multi-way repeated measures or mixed analyses of variance (ANOVAs), using CLEAVE software (Herron, 2005). The reported statistical values were corrected for potential violations of sphericity using the Greenhouse–Geisser epsilon.

#### III. RESULTS

Three analyses are presented in Secs. III A–III C. The first analysis focuses on the spatial profile of CP and its changeover time (Sec. III A). Then, the temporal profile of the CP is analyzed on time scales of minutes (Sec. III B) and seconds (Sec. III C).

# A. Spatial and temporal profiles of contextual plasticity

The upper panels of Fig. 2 show the across-subject mean bias in localization responses as a function of target location, separately for the two adaptor locations (circles for frontal vs triangles for lateral), the two context conditions (red for simple, blue for complex), and the two experiments (panels A and B for experiment 1; panel C for experiment 2). Panel B shows the experiment 1 data for the four subjects who also participated in experiment 2.

Two ANOVAs were performed on the bias data. The first ANOVA considered the experiment 1 data on all seven subjects (panel A), with factors of *adaptor location* (frontal, lateral), target location (seven locations from 11° to 79°), context type (simple, complex), and subrun (1-4). This ANOVA found significant main effects of adaptor location  $[F(1,6) = 56.93, p = 0.003, \eta_p^2 = 0.905]$  and *target location*  $[F(96,36) = 4.76, p = 0.0012, \eta_p^2 = 0.442]$ , as well as significant interactions of *target location*  $\times$  *subrun* [F(18,108) =2.24, p=0.0057,  $\eta_p^2$ =0.272], adaptor location × subrun  $[F(3,18) = 27.52, p < 0.0001, \eta_p^2 = 0.821], context type$ × adaptor location [F(1,6) =11.49, p = 0.0147,  $\eta_p^2 = 0.657$ ], and context type  $\times$  target location [F(6,36) = 3.95, p = 0.0039,  $\eta_p^2 = 0.397$ ]. The second ANOVA considered both rooms and was restricted to the four subjects who performed the experiments in both rooms (panels B and C). It had an additional factor of *room* (anechoic, reverberant), and it found significant interactions of *context type*  $\times$  *room*  $\times$  subrun  $\times$  adaptor location [F(3,9) = 4.84, p = 0.0454,  $\eta_p^2 = 0.617$ ), context type  $\times$  target location [F(6,18) = 6.19, p = 0.0387,  $\eta_p^2 = 0.674$ ], context type  $\times$  adaptor location  $[F(1,3) = 12.54, p = 0.0383, \eta_p^2 = 0.807], \text{ and } subrun$ × adaptor location [F(3,9) = 35.62, p = 0.0005,  $\eta_p^2 = 0.922$ ].





FIG. 2. (Color online) Upper panels show mean response biases ( $\pm$  standard error of the mean) in target-only trials in experiment 1 (panel A) and experiment 2 (panel C), plotted as a function of target location separately for each combination of context type and adaptor location. Panel B shows the experiment 1 data for the four subjects who participated in experiment 2. For each panel in the upper row, a pair of panels in the lower row shows the buildup of biases during subruns, separately averaged across the group of targets at azimuths of  $11^{\circ}-34^{\circ}$  (left-hand panel of each pair) and at azimuths  $56^{\circ}-79^{\circ}$  (right-hand panel of each pair). Lat Adaptor, lateral adaptor.

Finally, it found a significant main effect of *adaptor location* [F(1,3) = 2065.76, p < 0.0001,  $\eta_p^2 = 0.998$ ], while no other main effects or interactions reached significance.

The data in the upper row of Fig. 2 indicate that localization responses were biased relative to the actual target locations. The frontal context data (circles) were biased laterally by approximately  $5^{\circ}-13^{\circ}$ , while the lateral context data (triangles) were biased by -5 to  $5^{\circ}$ . Such "global" response biases are common in localization experiments and arise from a combination of factors, including the response method (Kopčo *et al.*, 2015). Since we did not measure a no-context baseline, the analysis here is focused on differences in the bias, depending on the context, as this difference was not expected to be influenced by factors like response method. The clearest effect shown in the upper panels of Fig. 2 is that the responses with frontal contexts are always biased more laterally than the responses with lateral contexts (triangles fall under circles in all three panels, confirmed by the main effect of *adaptor location*). This effect is overall stronger for the complex context than the simple context, particularly near the adaptor locations (blue circles are above the red circles, especially for the targets at  $11^{\circ}-34^{\circ}$ ; blue triangles are below the red triangles, especially for the targets at  $56^{\circ}-79^{\circ}$ ; significant *context type* × *adaptor location* and *context type* × *target location* interactions). Because this pattern is approximately symmetric and complementary (dominated by the frontal adaptor for frontal targets and the lateral adaptor for lateral targets), the differences between frontal and lateral adaptor contexts are approximately target location–independent (corresponding red lines are approximately parallel, as are the corresponding blue lines; *context type* × *target location* × *adaptor location* interaction is not significant).

Before comparing the results across the rooms, note that the results in panels A and B are very similar, i.e., that the subgroup of participants who also participated in



experiment 2 is representative of the larger group. Panels B and C show that the effect of context was also modulated by the room in which the stimuli were presented, and the ANOVA further suggests that the room effect changed over time (four-way *context type*  $\times$  *room*  $\times$  *subrun*  $\times$  *adaptor location* and two-way *subrun*  $\times$  *adaptor location* interactions). These interactions did not include the *target location* factor, suggesting again that the important features of CP are approximately target location–independent.

For each panel in the upper row of Fig. 2, two corresponding panels in the lower row show the temporal profile of adaptation, separately averaged across the targets at azimuths of 11°-34° (left-hand panel of each pair) and at azimuths of 56°-79° (right-hand panel of each pair). By this rearrangement, the temporal resolution of data could be increased threefold, as each of the original four subruns now contained three data points, approximately evenly distributed across it. These temporal data show two main results. First, even though they are fairly noisy, they provide a rough idea about where the pre-adaptation, no-context baseline responses would have fallen, if they were measured, as the difference between frontal and lateral context effects are always the smallest at the beginning of the run. So, for example, it can be estimated that the baseline bias in responses to targets 11°-34° was approximately 8°-10°, as this is the average of the circles and squares at the beginning of the run for these targets in all three lower left sub-panels of panels A, B, and C (and, in particular, for the simple context in experiment 1, the circles and triangles initially overlap, suggesting no contextual adaptation influenced these data yet). For targets at  $56^{\circ}$ – $79^{\circ}$ , the baseline values can be estimated to be lower, approximately  $0^{\circ}$ – $5^{\circ}$  (average of the initial values in lower right sub-panels of panels A, B, and C). Second, these temporal data show that the effect of contextual adaptation was always a repulsion, as all the frontal context profiles grow more lateral over time (circles in all panels have an increasing trend), while the lateral context profiles grow more frontal (triangles in all panels have a decreasing trend).

We operationalize CP in terms of CP<sub>diff</sub>, which is the difference between frontal and lateral context biases, averaged across all seven target locations. CP<sub>diff</sub> is plotted as a function of subrun in Fig. 3(A) for both experiments (differentiated by line styles, which match Fig. 2) and contexts (line color).<sup>1</sup> The results show that, overall, CP had a fast onset, reaching values between 3° and 7° within the first subrun. It continued to grow on the time scale of minutes in both conditions [top scale in Fig. 3(A)], with the rate of growth dependent on the context type and on the room. Overall, CP tended to be larger with the complex context (blue lines are above red lines) and in the anechoic room (dashed lines tend to be above solid lines). However, these effects varied over time and did not combine additively. Specifically, for the simple context (red lines), the room effect (dashed vs solid line) was the largest at the beginning of the run, while for the complex context (blue lines), it was largest at the end of the run. Finally, the complex context CP in both environments continued to grow between the



FIG. 3. (Color online) Temporal profile of  $CP_{diff}$  for the two contexts and two rooms. (A) Mean  $CP_{diff}$  ( $\pm$  standard error of the mean) divided into subruns and rearranged to increase temporal resolution. (B) Simple context data modeled using exponential fits (fitted parameters are shown in the inset). (C) Complex context re. simple context data modeled using linear fits (parameters shown in the inset).

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third and fourth subruns, suggesting that it might not have reached its maximum over the 5 min course of individual runs in this condition (reaching  $12^{\circ}-14^{\circ}$ ). In the simple context condition, CP appeared to reach its maximum of  $8^{\circ}-10^{\circ}$  by the third subrun.

# 1. Discussion

The analysis of the spatial properties of CP showed that (1) CP<sub>diff</sub> is observed as a repulsion of responses away from the adaptor location that decreases with separation between target and adaptor, and that (2) the effect is stronger in the complex context condition where the overall adaptor click rate is higher. These results are consistent with previous studies (Hládek *et al.*, 2017; Kopčo *et al.*, 2015) but extend the finding to lateral, as well as frontal, adaptors. Additionally, we find that context effects are similar both in their strength and spatial extent for the frontal and lateral adaptor, suggesting that the spatial representation in which CP is induced is approximately uniform, even though auditory spatial resolution decreases with azimuth (Hartmann and Rakerd, 1989).

The temporal analysis of trends for different target regions (Fig. 2, bottom panels) showed that there were likely lateral biases in responses independent of the contextual effects, visible at the beginning of the runs, and that these biases were larger for the frontal targets than the lateral targets. It also showed that, during the runs, the effects of context expressed as  $CP_{diff}$  [Fig. 3(A)] were always repulsive (away from the adaptor location). This result is consistent with the studies by Kopčo *et al.* (2015) and Hládek *et al.* (2017), which did measure the no-context baseline.

The temporal analysis of CP showed that the effects of room and context type interact and are non-additive. Specifically, CP was strong already at the beginning of the run; in both rooms for the complex context and in the anechoic room also for the simple context, while being relatively weak in the classroom simple context runs. Toward the end of the runs, CP became largely independent of the environment while differing strongly for the two contexts. Specifically, the complex context CP continued to grow for as long as 5 min, while in the simple context, the CP reached its maximum after 2-3 min, consistent with previous studies that only used simple context (Hládek et al., 2017). Thus, varying the context from trial to trial causes at least a doubling of the time it takes CP to reach its asymptote. It also appears that varying the context results in CP that is stronger  $(12^{\circ}-14^{\circ})$  by subrun four) than that reported with a simple eight-click context (CP of up to 9°) (Hládek et al., 2017) and, thus, that the variation might be important. However, note that Hládek et al. (2017) reported CP relative to a nocontext baseline, not CP<sub>diff</sub>, which might account for the apparent difference in strength. Thus, the number of clicks in the adaptor might still be the main determinant of the CP strength.

# B. Modeling of the temporal profile of contextual plasticity

We used exponential fits to analyze the buildup of CP in the simple context runs, and linear fits to describe the additional buildup in the complex context runs. Specifically, each subject's simple context CP data were fitted parametrically using the first-order exponential equation

$$y(t) = y_{\infty}(1 - e^{-t/\tau}),$$
 (1)

with time t in the units of subruns, yielding a time constant  $\tau$ for the adaptation to the context (with  $1/\tau$  as its rate), and a model estimate of the asymptotic value of CP,  $y_{\infty}$ . The model assumed that the initial, pre-adaptation CP was 0 and that the asymptotic value of CP was equal for the two environments [as suggested by the data in Fig. 3(B)]. Thus, only three parameters were fitted per subject, one  $y_{\infty}$  for both environments and one  $\tau$  for each environment. The additional CP observed in the complex context (vs simple context) was modeled using a linear model as there was no evidence that the difference data deviated from linearity in either environment [Fig. 3(C)]. The analysis was then focused on the estimated slope of the adaptation, which represents the temporal properties of the additional adaptation. For each model, the goodness of fit is reported as the average of the coefficient of determination  $r^2$  computed for individual subject fits.

The results of this analysis are shown in Figs. 3(B) and 3(C). In both figures, the mean data are shown by symbols "x," the fits for the classroom are shown by thick solid lines (the four-subject fit is shown by thin solid lines), and the fits for the anechoic room are shown by dotted lines.

The simple context data and fits in Fig. 3(B) (red symbols and lines, corresponding to the red lines from panel A) show that the onset of CP is faster in the anechoic room than in the classroom, and that the difference between the anechoic and classroom data is only around 1° toward the end of the run. As mentioned above, given the small difference at the asymptote, the model was fitted such that only one common  $y_{\infty}$  was used for both environments while  $\tau$ values were separate. The common fitted value of  $y_{\infty}$  was 10.2°. On the other hand, the time constant  $\tau$  differed significantly between the environments for the four subjects who participated in both experiments. The mean  $\tau$  was 2.3 subruns in the classroom and 0.6 subruns in the anechoic room [t(3) = -3.18, p = 0.02]. The across-subject average  $r^2$  of the models is 0.32 in the classroom (0.48 for four subjects) and 0.22 in the anechoic room.

Figure 3(C) shows, for both environments, the difference between the complex and simple context CP data (i.e., the difference between respective blue and red data from panel A), as well as the linear fits. The data show that the effect of complex context (re. simple context) is approximately linear in both environments. In the classroom experiment, the complex context caused an additional repulsion from the adaptor location of approximately  $4^\circ$ , independent of time (solid lines). In the anechoic room, the effect of the

complex context was much slower, growing from approximately  $0^{\circ}$  to  $4^{\circ}$ . A paired-samples t-test [t(3) = -4.7, p = 0.018] confirmed that the slopes of the fits were significantly different in the anechoic vs reverberant room. The across-subject average  $r^2$  of the models is relatively low, 0.01 in the classroom (0.11 for four subjects) and 0.1 in the anechoic room.

# 1. Discussion

The modeling presented in Sec. III B allowed us to quantify and extend the results of the behavioral data analysis of Sec. III A. Specifically, the exponential model fitted to the simple context data found a significant difference in adaptation rate between the two rooms, supporting the conclusion that the initial difference between the two rooms was mostly driven by a difference in speed, not strength, of CP, as the rate of change in the anechoic room was at least twice that of the classroom.

The additional adaptation in the complex context showed either a constant or an approximately linear growth, uniform over the whole duration of the current runs. This again shows that the complex context, randomly switching between a one-click and an eight-click adaptor, causes the adaptation to have an additional very slow component, much slower than that observed in our previous studies (e.g., Hládek *et al.*, 2017) and resulting in at least as strong, or stronger, CP. Note that the additional adaptation would likely have reached an asymptotic value if the runs were sufficiently long. However, since it did not reach its asymptote in the current experiment, and since the difference between the conditions was largely linear in both environments, a linear model was sufficient to describe the data.

Finally, note that the presented modeling always considered the difference between the frontal adaptor and lateral adaptor contexts, i.e., the  $CP_{diff}$ , corresponding to a combination of two adaptive processes, one for each context. The Appendix provides the results of additional modeling performed separately for the two adaptor locations, which shows that the slow minute-scale adaptation correlates with the distribution of the stimuli in different contexts, consistent with the hypothesis that spatial auditory processing prioritizes discriminability of stimuli over localization accuracy (Dahmen *et al.*, 2010; Lingner *et al.*, 2018).

#### C. Trial-to-trial adaptation in the complex context runs

Section III B showed that one effect of varying the context on a trial-to-trial basis was that the adaptation continued to evolve over the duration of an experimental run (around 5 min). Here, an analysis is performed on the time scale of individual trials, to examine (1) whether the extremely slow adaptation is accompanied and/or caused by a fast-varying plasticity changing after every context trial, and (2) whether this effect varies over the course of an experimental run. The complex context runs included two types of context trials (one-click and eight-click adaptors), randomly interleaved with the target-only trials. In this analysis, the target-only trials were split by the type of the preceding context trial (one-click or eight-click) and plotted as a function of subrun for the classroom [Fig. 4(A)] and the anechoic room [Fig. 4(B)]. In the classroom, the trials preceded by an eight-click adaptor trial show a faster onset of CP, reaching the maximum of  $12^{\circ}$  by the second subrun (solid lines with asterisks), while the trials preceded by a one-click adaptor trial show CP of around  $6^{\circ}$  in the first two subruns and reach  $12^{\circ}$  adaptation only in the third subrun (lines with no symbols). On the other hand, in the anechoic room, CP grows throughout the run with no systematic difference based on the immediately preceding context trial type.

Confirming these observations, an ANOVA with the factors of *subrun* (1–4) and *context* (one-click vs eight-click) performed on the classroom data found a significant main effect of *subrun* [F(3,18) = 14.93, p = 0.0003,  $\eta_p^2 = 0.71$ ] and a significant interaction *context* × *subrun* [F(3,18) = 3.59, p = 0.034,  $\eta_p^2 = 0.37$ )]. A similar ANOVA performed on the anechoic data found no significant effects or interactions [the main effect of *subrun* was no longer significant after the Geisser–Greenhouse correction [F(3,9) = 6.6, p = 0.063].

## 1. Discussion

This analysis shows that, in some instances, CP was affected by the immediately preceding context trial, such that the effect was larger following eight-click vs one-click adaptors. This suggests a relatively fast contextual effect, corresponding to the 3-5 s time scale of individual trials. However, it was only observed during the first half of each run and only in the classroom. One possible explanation of the lack of an effect in the anechoic room is that, in addition to a slow adaptation, a reverberation–suppression mechanism related to precedence buildup (Brown *et al.*, 2015)



FIG. 4. (Color online) Effect of the context adaptor click rate (one-click vs eight-click adaptor) in the immediately preceding trial on CP in the complex context condition.  $CP_{diff}$  is plotted as a function of subrun in the class-room (panel A) and the anechoic room (panel B). Front. adapt., frontal adaptor; Lat. Adapt., lateral adaptor; Subjs, subjects.

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influenced these trials. Specifically, the repeated presentation of clicks in the eight-click adaptor may activate a reverberation-suppression mechanism that suppresses the spatial percept of the target on the subsequent trial along with the reflections from the adaptor clicks. This suppression might increase the subject's uncertainty about the target location, which might result in biases further away from the adaptor (e.g., since the adaptor location is a priori known and the subject will not respond past it, even if that is the perceived location due to uncertainty), and thus, an increased CP. When only one adaptor click is presented in the context, this suppression might be reduced as it does not get a chance to build up. It is not clear why this effect was restricted to the early part of each classroom run, though it is possible that the effect was simply not visible later as the CP saturated. An alternative mechanism might be related to perceptual organization, as the eight-click adaptor trials were designed to increase the perceptual segregation of the adaptors and targets (Kopčo et al., 2017). In any case, our data clearly show that each localization trial can be influenced by the immediately preceding trials, which may be an issue for task designs that intermix different conditions (e.g., Kopčo et al., 2017; Moore et al., 2020).

# **IV. GENERAL DISCUSSION**

Our work shows that the spatial and temporal distribution of stimuli (or "context") in which listeners perform a localization task has a complex influence on their behavior. The main finding of the current study is that the repeated presentation of an adaptor-target context induced a slow adaptation in the localization of targets that (1) resulted in biases of up to 14° away from the adaptor location, (2) built up over at least 5 min, and (3) depended on the spatial and temporal structure of the adaptors, as well as on the presence of reverberation. Specifically, increasing the average number of adaptor clicks (complex context) resulted in a stronger CP, while both varying the number of clicks from trial to trial and an exposure to reverberation resulted in a slower temporal profile of the adaptation (compared to the simple context in the anechoic room). Strikingly, the complex context resulted in adaptation that grew over time in both environments, resulting in the strongest CP we have observed to date. These effects of context type and environment are likely due to some non-linear interaction of multiple adaptive processes that depend on the adaptor presentation rate, its variability, the presence of reflections, and possible other, so far unexplored factors (e.g., presentation level).

The spatial profile of CP was originally reported to be largely independent of the adaptor-target distance (Kopčo *et al.*, 2007). Later studies, which only used a frontal adaptor and also included no-adaptor baseline runs, showed that the effect is stronger near the adaptor location and that it largely disappears for targets separated by  $80^{\circ}$  from the adaptor (Hládek *et al.*, 2017; Kopčo *et al.*, 2015). The current study showed that the dependence of the CP strength on the separation from the adaptor also applies to the lateral adaptors, and that the repulsive effects of frontal and lateral adaptors are similar. It is worth noting, however, that the adaptor was always at the edge of the target range in the current study. It is possible that placing the adaptor in the middle of the target range and/or using targets symmetrically located around the midline, as in the previous localization aftereffect studies, (Carlile *et al.*, 2001; Laback, 2023; Phillips and Hall, 2001; Thurlow and Jack, 1973) would result in a different pattern of adaptation.

The main finding concerning the temporal profile was that varying the context from trial to trial produced CP that was very slow to stabilize, continuing to grow for at least 5 min, while simple context CP asymptoted after 1-2 min (Hládek et al., 2017; Kopčo et al., 2015). Such extended adaptation has not been reported in previous localization aftereffect studies, which focused on effects occurring immediately post-adaptor and not analyzing the time course of buildup (Carlile et al., 2001; Lingner et al., 2018), while other related studies likely observed such long-term adaptation but ascribed it to other factors (Moore et al., 2020). In future studies, it would be very interesting to include long enough runs for CP to reach the asymptotic value in the complex context, so that it can be established how long such an adaptation can continue. Another finding was that, in the complex context, there was evidence for a fast adaptive component that is sensitive to the temporal structure of individual context trials. Since this fast component was not observed in the anechoic room, it is possible that it is related to reverberation-suppression mechanisms evoked in the precedence effect and its buildup (Brown et al., 2012; Litovsky and Macmillan, 1994) or perceptual organization (Kopčo et al., 2017). On the other hand, our subject pool may have been too small to reveal similar effects in the anechoic room, and further investigations would be needed to make strong conclusions. Reverberation also affected the initial onset of CP, which was considerably slower in the classroom than in the anechoic room for the simple context. Again, this difference may be related to precedence buildup mechanisms operating in the reverberant classroom. Overall, the effects of reverberation that we observed were small, and we did not find strong support for the hypothesis that CP would be weaker in reverberation where the presence of omnidirectional reverberation makes the distribution of energy more uniform around the listener.

The stimuli used in our studies of CP differ in several aspects from previous localization aftereffect studies (Carlile *et al.*, 2001; Dingle *et al.*, 2012; Laback, 2023; Phillips and Hall, 2005). First, while in CP, the adaptors are brief 2 ms clicks (or eight-click trains), previous studies used adaptors ranging in duration from 600 ms to minutes, meaning that the cumulative duration of the adaptor stimulus within a whole run of the current study is less than the adaptor duration in a single trial of the previous studies. Second, while in CP, the gap between the adaptor and target is around 3-5 s, in the previous studies, it was much shorter (from 1 s down to 10 ms). Third, the to-be-localized target

stimuli are typically much longer (hundreds of milliseconds) in localization after-effect studies than the CP clicks.

The characteristics of CP also differ in key ways to reported localization aftereffects. For example, we showed that CP built up over minutes and, in certain conditions, it was only detectable after a minute of adaptation. Most localization aftereffect studies in the literature did not report any temporal dynamics. One exception is the recent study by Laback (2023) who looked for possible "across-trial" precursor effects but did not find evidence of a slow buildup like that reported here. Thus, while it is possible that CP is caused by the same adaptation mechanisms as the localization aftereffect, CP demonstrates the dynamics of this mechanism when excited by much briefer and sparser adaptors. Moreover, CP is clearly sensitive to reverberation, suggesting that the reverberation-suppression mechanism might also contribute to CP. Finally, CP is sensitive to the similarity between adaptor and target (Hládek et al., 2017) and only has been reported in studies in which the subject is involved in active localization on the contextual trials (Kopčo et al., 2015). Thus, perceptual organization and changes in localization strategy might also affect CP. Overall, there is still much to explore in terms of the relationship between CP and the localization aftereffect, and the extent to which they are driven by shared mechanisms.

Finally, while the data presented here are unable to distinguish between competing models of spatial adaptation that have been proposed in the literature, they provide some preliminary indications that may be worth following up on. For example, an exploratory analysis of the data reported in the Appendix shows that the slow component of CP can be well characterized as a linear drift in the spatial auditory representation in response to the overall spatial distribution of the stimuli in a particular run. Specifically, stronger drifts toward midline were observed with increased laterality of the distribution mean. Such a relationship is consistent with the idea that CP might be caused by adaptation of the neural representation that shifts it toward the stimulus distribution mean (Dahmen et al., 2010; Lingner et al., 2018). The specific neural mechanisms underlying the shift might include dynamic range adaptation (Dahmen et al., 2010), synaptic gain control (Stange et al., 2013), or re-balancing of excitatory and inhibitory inputs (Magnusson et al., 2008). This "shift" model offers an alternative to "suppression" models which posit that localization aftereffects are caused by local suppression/fatiguing of spatial channels near the adaptor (e.g., Carlile et al., 2001). Of course, it is possible that both shift and suppression mechanisms contribute to CP and related spatial adaptation phenomena. Finally, the mechanisms of precedence buildup (Clifton and Freyman, 1997), and those proposed in the three-channel model of adaptation (Dingle et al., 2012), might also be contributing. Future experiments specifically designed to untangle these mechanisms may bring further insights.

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#### **APPENDIX**

#### Relationship between contextual plasticity buildup and stimulus distribution

#### Motivation

In this exploratory analysis, we attempted to relate the temporal profile of CP to the spatial distribution of the stimuli in different contexts. Our goal here was to provide a preliminary test of competing hypotheses about the mechanisms underlying CP.

While the mechanism underlying CP is largely unknown, it shares many properties with the localization aftereffect (Phillips and Hall, 2005; Thurlow and Jack, 1973). Specifically, it results in similar shifts in the perceived target location away from the adaptor location. Various models have previously been proposed for the localization aftereffect, many of them assuming that it is caused by some suppression in the neural representation of auditory space (Carlile et al., 2001; Dingle et al., 2012). It has also been suggested that the observed shifts are a result of a broad dynamic range adaptation of the auditory spatial representation, occurring when the stimulus distribution becomes concentrated in a subregion of the full horizontal spatial range (Dahmen et al., 2010). In this scenario, biases in responses are a negative side effect of the representation adapting to improve the spatial separability of targets presented within the subregion. This adaptation may be implemented by fitting the working point of the neural firing rates vs the spatial location to the middle of the stimulus range (Lingner et al., 2018).

Motivated by the latter studies, here we examine the hypothesis that the auditory representation adapts to the non-uniform stimulus distribution in our experiments. We test a simplified prediction that the more skewed the stimulus distribution from the midline, the stronger the response biases induced by it. To test this prediction, we analyze the drifts in response biases over the course of individual runs from subrun 1 to subrun 4 and evaluate whether the slope of these drifts, averaged across target location, can be predicted by the size of the change in the stimulus distribution mean. The analysis focuses on the drifts, not on the absolute value of the change, because looking at the drifts (1) allows us to consider the frontal context and lateral context data separately, as we are only looking within a run, (2) only requires us to use the first subrun as a reference (no preadaptation reference was measured), and (3) allows the analysis to focus on the slow adaptation occurring on a time scale larger than  $1-2\min$  (i.e., the approximate duration of one subrun),

in which the drifts were largely linear. Our analysis is performed on the data presented in the main body of the current paper, as well as on additional data from Kopčo *et al.* (2015).

#### Data from current study

In experiments 1 and 2, targets were presented from a frontal left or frontal right quadrant in the horizontal plane, with the adaptor always located at the edge of the target range (Fig. 1). The left-hand panel of Fig. 5(A) shows the distributions of these click stimuli within a run, including both the adaptor and target clicks (bars), separately for the frontal adaptor and lateral adaptor runs (note that the distribution was identical in the two experiments). The symbols along the upper edge indicate the respective distribution means. In each of the four contexts, the stimuli are shown for the runs performed in the right-hand quadrant (the lefthand quadrant stimuli would add flipped distributions and means on the left-hand side). The distribution mean was between 9° for the frontal adaptor, complex context runs (blue circle), and 81° for the lateral adaptor complex context runs (blue triangle), with the respective simple context means (red circle and triangle) falling between the complex context mean values. Based on our hypothesis, for the stimuli presented in the right-hand quadrant, the responses are expected to drift to the left, as the channels representing the left quadrant shift their receptive fields to the right. Additionally, this drift is expected to be larger in the lateral adaptor runs than in the frontal adaptor runs, as the distribution is skewed more positively (to the side) when the adaptor is lateral.

The two left-hand panels of Fig. 5(B) show, for the two experiments, the across-target average response bias as a function of subrun, considering only the four subjects who participated in both experiments. The symbols represent the mean response bias in each subrun, separately for the simple vs complex contexts (red vs blue), frontal vs lateral adaptors (circles vs triangles), and classroom vs anechoic room (filled vs open symbols). The lines show the across-subject average of linear fits through the data performed separately for each context (lines going through the triangle vs circle data for the frontal vs lateral context) and room (solid for classroom, dashed for anechoic room). The fits are very good, indicating that the adaptation is approximately linear over this time range. They have negative slopes for the lateral context (from -1.8 to  $-0.94^{\circ}$ /subrun) and slightly positive slopes for the frontal context (0.06°-0.38°/subrun), an effect that tends to be stronger for the complex vs simple context (blue vs red lines), especially in the anechoic room (dashed lines). These trends were confirmed by an ANOVA with factors of context type, room, and adaptor location, performed on the slopes of the linear fits, which found a significant main effect of *adaptor location* [F(1,3) = 105.87, p = 0.002,  $\eta_p^2 = 0.9724$ ] and a significant three-way interaction [F(1,3) = 10.43, p = 0.048, $\eta_{\rm p}^2 = 0.7766$ ].





FIG. 5. (Color online) Relationship between stimulus distribution and drift in responses for current experiments 1 and 2 (left-hand panels) and from a previous study [(Kopčo *et al.*, 2015); right-hand panels]. (A) Bars show the distribution of click stimuli in experimental runs (considering both adaptor and target clicks) separately for each context. Symbols along the upper edge indicate the stimulus distribution mean. (B) For each context, symbols represent the across-subject mean of bias in responses averaged across target locations as a function of subrun. Lines represent corresponding linear fits, i.e., temporal drifts in the responses. (C) Symbols represent acrosssubject mean slope of the linear fit (from panel B) as a function of the stimulus distribution mean (from panel A), shown separately for each adaptor location (runs with frontal, lateral adaptor, or no adaptor). Lines show a linear fit of this relationship. Error bars represent standard error of the mean. Anech., anechoic; Base., baseline; Sim, simple; Comm., complex; Distrib., distribution; Fr., frontal; Front., frontal; Lat., lateral; Prob., probability.

To directly evaluate the relationship between the distribution of the stimuli and the response drifts, the left-hand panel of Fig. 5(C) plots the slope of the response drifts [from Fig. 5(B)] as a function of the mean lateral position of the stimuli [from Fig. 5(A)]. There is a strong correlation, with the across-subject average r reaching 0.95 in the



anechoic room and 0.86 in the classroom (the slightly weaker correlation in the classroom is mainly caused by the frontal data shown in blue vs red circles, which do not show the downward trend of all the other data). A linear fit to the data (black line) shows that the slope of the drift in responses is inversely proportional to the mean of the stimulus distribution [slope of this fit is -0.033; t(6) = -10.4, p < 0.0001]. This general result is consistent with the idea that the drift occurs as a result of a dynamic range adjustment (Dahmen et al., 2010; Lingner et al., 2018). However, there is one aspect of the data that is not consistent with this idea. While the distribution means are all positive, predicting that the drift slopes should always be negative, the slopes of the drift for the frontal contexts are slightly positive [circles in Fig. 5(C)]. A potential explanation for this discrepancy is that, in addition to the distribution-dependent drifts, the responses also drifted due to some other factors, like a fatiguing of the motor responses, as the subjects used a hand-held pointer to respond.

## Data from Kopčo et al. 2015

To examine whether the slopes are influenced by the response method used by the listeners, we performed the same analysis on data from a previous study (experiment 1 of Kopčo et al., 2015). That study was very similar to the current simple context classroom experiment 1, differing only in two important aspects. First, three different response methods were used: (1) using a hand-held pointer while the eyes were closed (like in the current study), (2) using a hand-held pointer with the eyes open, and (3) a keyboardbased method that used vision but did not require any sensory-motor spatial transformation to respond. The righthand panel of Fig. 5(A) shows the stimulus distribution in this study. The frontal adaptor runs had distributions very similar to experiment 1 (green vs red filled bars), while the baseline runs had a uniform distribution with a mean at 45° (black bars).

The right-hand panel of Fig. 5(B) shows the buildup of adaptation in response bias as a function of subrun, in a format similar to the left-hand panel. Here, the circles represent the frontal adaptor data and crosses the no-adaptor baseline data for all three response methods. The lines of different styles represent the linear fits for the different response methods, separately for the frontal adaptor (green) and baseline (black) runs. There are clear differences between the lines for the different response methods, both in terms of their mean values and their drifts (e.g., solid lines are the most positive and decreasing, whereas the dashed-dotted lines are the most negative and increasing). This confirms that a part of the drifts observed in experiments 1 and 2 might be due to drifts in responses, not due to adaptation in the auditory spatial representations. However, important for the current study, the differences between frontal and baseline lines corresponding to the same response method always have a similar pattern, with the former having a more positive slope than the latter (e.g., compare the green and black dashed-dotted lines). Thus, the relative change in the slope of the drift appears to be independent of the response methods and thus, may be related to adaptation in a spatial map. These results were confirmed by an ANOVA performed on the slope values with the factors of *response method* (three levels) and *adaptor location* (frontal, baseline) which found significant main effects of *response method* [F(2, 18) = 12.07, p < 0.001] and *adaptor location* [F(1,9) = 31.3, p < 0.0005] but no significant interaction.

The right-hand panel of Fig. 5(C) shows the relationship between the slope of the response drifts and the mean stimulus position for the three response methods (thin lines with different styles and the corresponding "o" and "x" symbols), as well as for their average (thick solid line). Consistent with the current experimental results, the average fit shows that the slope of the drift in CP is inversely proportional to the mean of the stimulus distribution [slope of this fit is -0.022; t(9) = -7.46, p < 0.0001]. Importantly, the large vertical offsets between the lines corresponding to the different response methods show that the drift slopes are responsemethod dependent. Thus, only the relative differences obtained with the same response method (or the slopes) can be ascribed to adaptation in the spatial representation.

# Discussion

This analysis showed that the slow drift in response bias is proportional to the mean lateral position of the stimuli, independent of potential drifts in motor responses or of the environment. Specifically, stronger drifts toward midline were observed with increased laterality of the distribution mean, consistent with the idea that CP might be caused by adaptation of the neural representation to the stimulus distribution such that the neural operating points or spatial channels shift toward the stimulus distribution mean (Dahmen et al., 2010; Lingner et al., 2018). This is an alternative to a previously proposed model suggesting that repulsion-byadaptor localization aftereffects might be caused by local suppression/fatiguing of the spatial neural channels near the adaptor caused by their extended stimulation (Carlile et al., 2001). While the current results are qualitatively consistent with such a suppression mechanism, as the responses also drift from the adaptors, the suppression mechanism does not predict that these drifts would grow with the adaptor laterality, as the suppression predicts the same amount of adaptation independent of the adaptor location (Fig. 6 in Carlile et al., 2001).

Note that the current analysis has several limitations. First, it assumes that the stimulus distribution mean is a relevant characterization of the distribution. Previous studies showed that other distribution statistics, like the standard deviation, also influence spatial adaptation (e.g., Dahmen *et al.*, 2010; Laback, 2023). Second, it only looks at the across-target mean drift in the responses, ignoring the fact that responses for some target locations might have drifted more than others. Future studies are needed to look at both other candidate statistics (e.g., stimulus variance, range,



distribution median, etc.) and on the dependence of the drifts on the target location.

- <sup>1</sup>Note that compared to the lower panels of Fig. 2, the error bars in Fig. 3(A) are much smaller, showing that there were overall individual biases in responses that are subtracted out when  $CP_{diff}$  is computed.
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